

# An Operational Measure of Information Leakage

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**Abstract**—Given two discrete random variables  $X$  and  $Y$ , an operational approach is undertaken to quantify the “leakage” of information from  $X$  to  $Y$ . The resulting measure  $\mathcal{L}(X \rightarrow Y)$  is called *maximal leakage*, and is defined as the multiplicative increase, upon observing  $Y$ , of the probability of correctly guessing a randomized function of  $X$ , maximized over all such randomized functions. It is shown to be equal to the Sibson mutual information of order infinity, giving the latter operational significance. Its resulting properties are consistent with an axiomatic view of a leakage measure; for example, it satisfies the data processing inequality, it is asymmetric, and it is additive over independent pairs of random variables. Moreover, it is shown that the definition is robust in several respects: allowing for several guesses or requiring the guess to be only within a certain distance of the true function value does not change the resulting measure.

**Index Terms**—Leakage, Privacy, Sibson mutual information, Inference

## I. INTRODUCTION

Given two discrete random variables  $X$  and  $Y$ , how much information does  $Y$  leak about  $X$ ? This basic question arises in many secrecy and privacy problems, in which  $X$  represents sensitive information and  $Y$  represents information available to an adversary. A quantitative answer to that question is necessary to assess the performance of privacy systems for which  $Y$  cannot be made independent of  $X$ , which is often the case in practice. For example, a curator might want to reveal statistical data about a given population without compromising the privacy of its individuals [1]–[3]. An adversary could also gain access to  $Y$  through a side channel [4]–[7], or through a wiretap [8,9]. Perfectly securing these channels, if even possible, could be highly detrimental to the performance of the underlying system.

Moreover, the question is interesting from a purely theoretical point of view, as it is akin to fundamental questions in information theory such as “how much information does  $X$  contain?” The fact that Shannon answered the latter question with the entropy of  $X$ ,  $H(X)$ , might explain why many works [8]–[15] have adopted equivocation,  $H(X|Y)$ , as a measure of privacy.

However, this choice overlooks the context in which these questions were posed. Whereas Shannon’s motivating problem was finding the minimum number of bits required to *describe*  $X$ , the goal in privacy problems is different. It also fails to capture the fact that the adversary might be interested in functions of  $X$ , or other random variables dependent on  $X$  [16,17], rather than  $X$  itself.

In this paper, we give an operational definition of leakage that is motivated by the setup of a *guessing* adversary. More specifically, upon observing  $Y$ , the adversary tries to guess a (possibly randomized) function of  $X$ . Leakage for a specific function is considered to be the logarithm of the ratio of the probability of a correct guess when  $Y$  is observed, to the probability of a correct guess when it is not (i.e., a blind guess). Maximal leakage, which we denote by  $\mathcal{L}(X \rightarrow Y)$ , is then defined as the maximum leakage over all such randomized functions. This maximization, which is formally over discrete random variables  $U$  for which the Markov chain  $U - X - Y$  holds, represents a worst-case analysis on the function of interest  $U$ , and models scenarios in which the conditional distribution  $P_{U|X}$  is unknown. It is also inspired by the strong data processing constant [18].

Although the maximization is an infinite-dimensional problem,  $\mathcal{L}(X \rightarrow Y)$  admits a closed-form solution. It turns out to equal the Sibson mutual information of order infinity  $I_\infty(X; Y)$  [19,20], endowing it with an operational significance. Several desirable properties for a leakage measure then follow: it is zero if and only if  $X$  and  $Y$  are independent, it is not symmetric, it satisfies the data processing inequality, and it is additive over independent pairs  $\{(X_i, Y_i)\}$ .

We provide a conditional probability law  $P_{U|X}$  that achieves the maximum and depends on the joint probability  $P_{XY}$  only through its  $X$ -marginal,  $P_X$ . In particular,  $P_{U|X}$  is such that: for distinct  $x$ ’s, the supports of  $P_{U|X=x}$ ’s are disjoint, and each  $P_{U|X=x}$  effectively “shatters” the atom  $x$  into (almost) uniformly distributed  $u$ ’s to get an (almost) uniform marginal  $P_U$ . Moreover,

we show that, in general, there is no deterministic law  $P_{U|X}$  that achieves the maximum. Indeed, we could have  $X$  and  $Y$  such that  $\mathcal{L}(X \rightarrow Y) > 0$ , whereas observing  $Y$  does not affect the probability of guessing any deterministic function of  $X$ .

Furthermore, we show that the definition of maximal leakage is robust in several respects. In the definition of  $\mathcal{L}(X \rightarrow Y)$ , we allow the adversary *one* guess only. A natural extension would be to allow for, say,  $k$  guesses for some integer  $k$ . This is particularly relevant for privacy problems. For example, if  $U$  is a password to some system, then an adversary is typically allowed several wrong guesses before he/she is possibly locked out. We call the modified measure  $k$ -maximal leakage, and denote it by  $\mathcal{L}^{(k)}(X \rightarrow Y)$ . We show that, in fact, the two definitions are equivalent for all  $k$ .

Finally, we consider the case in which the adversary only needs the guess to be within a certain distance of the true function value, according to an arbitrary distance metric. As such, the random variable  $U$ , over which we are optimizing, now lives in a given metric space  $\mathcal{U}$  and is no longer restricted to be discrete. We call this modified measure maximal locational leakage, and we denote it by  $\mathcal{L}_{\mathcal{U}}(X \rightarrow Y)$ . We show that  $\mathcal{L}_{\mathcal{U}}(X \rightarrow Y) \leq \mathcal{L}(X \rightarrow Y)$ , and equality holds under an unboundedness condition on the metric space  $\mathcal{U}$ .

## II. RELATED LEAKAGE METRICS

The literature on leakage and privacy measures is vast, spanning the fields of information theory, computer science, and computer security. The closest to our work comes from computer security in [21]–[24]. In particular, [21] defines leakage from  $X$  to  $Y$  as the logarithm of the multiplicative increase, upon observing  $Y$ , of the probability of guessing  $X$  *itself* correctly, neglecting that the adversary might be interested in certain functions of  $X$ . [22] considers a *worst case* approach, and maximizes the previous quantity over all distributions on the alphabet of  $X$  (while  $P_{Y|X}$  is fixed). The resulting quantity turns out to equal  $\mathcal{L}(X \rightarrow Y)$  and is called in the computer security literature “maximal leakage” as well, or min-capacity (the term min-capacity is slightly misleading as  $\mathcal{L}(X \rightarrow Y)$  is in fact greater than or equal to the Shannon capacity of the channel defined by  $P_{Y|X}$  [23]). However, the term “min” was used because the min-entropy appears when computing probabilities of correct guesses). It is denoted by  $ML(P_{Y|X})$ , and its properties were further studied in [23,24]. [25,26] investigate relationships between maximal leakage and differential privacy [27], which is the most widely adopted metric in database security. Roughly speaking, differential privacy requires

that, for any two *neighboring* databases, the probabilities of any given output do not differ significantly.

Another connected line of work stems from cryptography, and in particular from the notion of *semantic security* [28] which considers the security of encryption schemes. First, [28] introduces the notion of “advantage” for a given function of the messages. It is the *additive* increase, upon observing the encrypted message (i.e., the ciphertext), of the probability of correctly guessing the value of the function. In our framework, “advantage” is defined as the multiplicative increase. Since one is typically interested in securing hard-to-guess functions for which the probability of a correct guess is small (since, otherwise, we are already “doomed”), the multiplicative increase is arguably more descriptive of the change. It also makes more intuitive sense when viewing leakage in terms of leaked *bits*. Semantic security then requires that, for an adversary that can work only for a polynomial (in the length of the message) amount of time, the advantage is negligible for all *deterministic* functions that are computable in polynomial time, and for all input distributions.

There are several variants of semantic security. In particular, entropic security [29,30] drops the computational bounds (on the adversary and the considered functions), but restricts its attention to input distributions with high min-entropy. [31] introduces semantic security to the wiretap channel, and does not restrict it to computationally bounded adversaries, nor deterministic polynomial-time computable functions. For a given encryption scheme, [31] then upper and lower-bounds the advantage of semantic security in terms of – what the authors call – mutual information security advantage, which is defined as the maximum, over all input distributions, of the mutual information between the message and the output of the channel whose input is the encryption of the message. Moreover, for discrete random variables  $X$  and  $Y$ , [32] upper-bounds the advantage over all deterministic functions in terms of their maximal correlation, which inspired [33] to use the latter quantity as a secrecy metric. [32], inspired by the correspondence analysis literature [34], also generalized maximal correlation to  $k$ -correlation, which is defined as the sum of the  $k$  largest principal inertial components of the joint distribution  $P_{XY}$ .

Finally, rate-distortion-based approaches to privacy metrics can be found in [35]–[39]. Although the particular metrics differ among those works (e.g., expected distortion, probability of a guess satisfying the distortion constraint, etc.), they all assume that there is a known distortion function up to which the adversary is interested

in the sensitive information  $X$ . For further discussion of privacy metrics, we refer the reader to [40], which categorizes over eighty such metrics.

### III. MAXIMAL LEAKAGE

Let  $X$  and  $Y$  be two discrete random variables, with alphabets  $\mathcal{X}$  and  $\mathcal{Y}$  respectively. We denote by  $P_{XY}$  the joint distribution of  $(X, Y)$ .

*Definition 1 (Maximal Leakage):* Given a joint distribution  $P_{XY}$  on finite alphabets  $\mathcal{X}$  and  $\mathcal{Y}$ , the maximal leakage from  $X$  to  $Y$  is defined as

$$\mathcal{L}(X \rightarrow Y) = \sup_{U \sim X \rightarrow Y \rightarrow \hat{U}} \log \frac{\Pr(U = \hat{U})}{\max_{u \in \mathcal{U}} P_U(u)},$$

where  $U$  and  $\hat{U}$  take values in the same finite alphabet.

We can rewrite  $\mathcal{L}(X \rightarrow Y)$  as

$$\mathcal{L}(X \rightarrow Y) = \sup_{U \sim X \rightarrow Y} \log \frac{\sum_{y \in \mathcal{Y}} \max_{u \in \mathcal{U}} P_{UY}(u, y)}{\max_{u \in \mathcal{U}} P_U(u)}. \quad (1)$$

Our main theorem is the characterization of maximal leakage as follows.

**Theorem 1:** For any joint distribution  $P_{XY}$  on finite alphabets  $\mathcal{X}$  and  $\mathcal{Y}$ , the maximal leakage from  $X$  to  $Y$  is given by

$$\mathcal{L}(X \rightarrow Y) = \log \sum_{y \in \mathcal{Y}} \max_{\substack{x \in \mathcal{X}: \\ P_X(x) > 0}} P_{Y|X}(y|x).$$

*Proof:* Assume, without loss of generality, that  $P_X(x) > 0$  for all  $x \in \mathcal{X}$ . Note that the right-hand side is equal to  $I_\infty(X; Y)$  [19,20]. To show that  $\mathcal{L}(X \rightarrow Y) \leq I_\infty(X; Y)$ , consider any  $U$  satisfying  $U \sim X \rightarrow Y$ . Let

$$\mathcal{L}(X \rightarrow Y)[U] = \log \frac{\sum_{y \in \mathcal{Y}} \max_{u \in \mathcal{U}} P_{UY}(u, y)}{\max_{u \in \mathcal{U}} P_U(u)}, \quad (2)$$

so that  $\mathcal{L}(X \rightarrow Y) = \sup_{U \sim X \rightarrow Y} \mathcal{L}(X \rightarrow Y)[U]$ . Then,

$$\begin{aligned} & \sum_{y \in \mathcal{Y}} \max_{u \in \mathcal{U}} P_{UY}(u, y) \\ &= \sum_{y \in \mathcal{Y}} \max_{u \in \mathcal{U}} \sum_{x \in \mathcal{X}} P_X(x) P_{U|X}(u|x) P_{Y|X}(y|x) \\ &\leq \sum_{y \in \mathcal{Y}} \max_{u \in \mathcal{U}} \sum_{x \in \mathcal{X}} P_X(x) P_{U|X}(u|x) \max_{x' \in \mathcal{X}} P_{Y|X}(y|x') \\ &= \sum_{y \in \mathcal{Y}} \left( \max_{x' \in \mathcal{X}} P_{Y|X}(y|x') \right) \max_{u \in \mathcal{U}} \sum_{x \in \mathcal{X}} P_X(x) P_{U|X}(u|x) \\ &= \sum_{y \in \mathcal{Y}} \max_{x \in \mathcal{X}} P_{Y|X}(y|x) \max_{u \in \mathcal{U}} P_U(u). \end{aligned}$$

Therefore,  $\mathcal{L}(X \rightarrow Y)[U] \leq I_\infty(X; Y)$  for all  $P_{U|X}$ , hence  $\mathcal{L}(X \rightarrow Y) \leq I_\infty(X; Y)$ .

For the reverse inequality, we construct a  $P_{U|X}$  for which  $\mathcal{L}(X \rightarrow Y)[U] = I_\infty(X; Y)$ . To that end, let  $p^* = \min_{x \in \mathcal{X}} P_X(x)$ . For each  $x \in \mathcal{X}$ , let  $k(x) = P_X(x)/p^*$ , and let  $\mathcal{U} = \bigcup_{x \in \mathcal{X}} \{(x, 1), (x, 2), \dots, (x, \lceil k(x) \rceil)\}$ . For each  $u = (i_u, j_u) \in \mathcal{U}$  and  $x \in \mathcal{X}$ , let  $P_{U|X}(u|x)$  be:

$$P_{U|X}((i_u, j_u)|x) = \begin{cases} \frac{p^*}{P_X(x)}, & i_u = x, \quad 1 \leq j_u \leq \lceil k(x) \rceil, \\ 1 - \frac{(\lceil k(x) \rceil - 1)p^*}{P_X(x)}, & i_u = x, \quad j_u = \lceil k(x) \rceil, \\ 0, & i_u \neq x, \quad 1 \leq j_u \leq \lceil k(i_u) \rceil. \end{cases}$$

*Remark 1:* It is easy to check that if  $\lceil k(x) \rceil = \lceil k(x) \rceil$ , then the corresponding formulas are equal.

Then, for each  $((i_u, j_u), x) \in \mathcal{U} \times \mathcal{X}$ ,

$$P_{UX}((i_u, j_u), x) = \begin{cases} p^*, & i_u = x, \quad 1 \leq j_u \leq \lceil k(x) \rceil, \\ P_X(x) - (\lceil k(x) \rceil - 1)p^*, & i_u = x, \quad j_u = \lceil k(x) \rceil, \\ 0, & i_u \neq x, \quad 1 \leq j_u \leq \lceil k(i_u) \rceil. \end{cases}$$

As mentioned in the introduction, the supports of  $P_{U|X=x}$  are disjoint for distinct  $x$ 's, and each  $x$  is effectively shattered into shards of probability  $p^*$ . Now, note that

$$\max_{u \in \mathcal{U}} P_U(u) = \max_{(i_u, j_u) \in \mathcal{U}} P_{UX}((i_u, j_u), i_u) = p^*. \quad (3)$$

Now, consider any  $(u, y) \in \mathcal{U} \times \mathcal{Y}$ . We have

$$\begin{aligned} & P_{UY}((i_u, j_u), y) \\ &= \sum_{x \in \mathcal{X}} P_X(x) P_{U|X}((i_u, j_u)|x) P_{Y|X}(y|x) \\ &= P_X(i_u) P_{U|X}((i_u, j_u)|i_u) P_{Y|X}(y|i_u) \\ &= \begin{cases} p^* P_{Y|X}(y|i_u), & 1 \leq j_u \leq \lceil k(i_u) \rceil, \\ (P_X(x) - (\lceil k(x) \rceil - 1)p^*) P_{Y|X}(y|i_u), & j_u = \lceil k(i_u) \rceil. \end{cases} \end{aligned}$$

Then, for a given  $y \in \mathcal{Y}$ ,

$$\begin{aligned} \max_{(i_u, j_u) \in \mathcal{U}} P_{UY}((i_u, j_u), y) &= \max_{(i_u, 1) \in \mathcal{U}} p^* P_{Y|X}(y|i_u) \\ &= \max_{x \in \mathcal{X}} p^* P_{Y|X}(y|x). \end{aligned} \quad (4)$$

Finally, we get

$$\mathcal{L}(X \rightarrow Y) \geq \mathcal{L}(X \rightarrow Y)[U] = \log \sum_{y \in \mathcal{Y}} \max_{x \in \mathcal{X}} P_{Y|X}(y|x),$$

where the inequality follows from the definition, and the equality follows from equations (2), (3), and (4). ■

The above result and analysis sheds light on the reason behind the equivalence between  $\mathcal{L}(X \rightarrow Y)$  and  $ML(P_{Y|X})$ . First,  $\mathcal{L}(X \rightarrow Y)$  depends on  $P_X$  only through its support. Moreover, on the one hand, the maximizer for  $ML(P_{Y|X})$  is always the uniform

distribution [22]; and on the other hand, for a uniform  $P_X$ , the above optimizing  $P_{U|X}$  is simply the identity map, which is the function of interest in [21,22].

In light of this, one might wonder if there is always a deterministic map  $P_{U|X}$  that achieves  $\mathcal{L}(X \rightarrow Y)$ . This is, however, not true in general. Suppose  $P_{XY}$  satisfies the following condition: there exists  $x^* \in \mathcal{X}$  such that for all  $y \in \mathcal{Y}$ ,  $P_{X|Y}(x^*|y) \geq 1/2$ . Then, for any deterministic function  $f$ ,  $\mathcal{L}(X \rightarrow Y)[f(X)]$  (cf. (2)) is zero since  $f(x^*)$  is always the optimal choice for the adversary, with and without the observation of  $Y$ . The above condition, however, is not sufficient for  $X$  and  $Y$  to be independent. The equivalence between  $\mathcal{L}(X \rightarrow Y)$  and  $I_\infty(X; Y)$  implies, on the other hand, that the independence of  $X$  and  $Y$  is necessary for maximal leakage to be zero, (see Corollary 2). Due to its usefulness, we state this equivalence as a separate corollary.

**Corollary 1:** For any joint distribution  $P_{XY}$  on finite alphabets  $\mathcal{X}$  and  $\mathcal{Y}$ ,

$$\mathcal{L}(X \rightarrow Y) = I_\infty(X; Y),$$

where  $I_\infty(X; Y)$  is the Sibson mutual information of order infinity.

Sibson's  $I_\alpha(X; Y)$  ( $\alpha \geq 0$ ) is an extension of the concept of Renyi entropy  $H_\alpha(X)$  and Renyi divergence  $D_\alpha(P||Q)$ . Although there are other possible extensions, [19] argues for the adoptions of Sibson's definition. Our result could be seen as also supporting that claim (more recently,  $I_\infty(X; Y)$  has been used as a complexity measure in the study of communication complexity [41]).

For binary-valued  $X$ , say  $\mathcal{X} = \{0, 1\}$ , Sibson [20] showed that

$$I_\infty(X; Y) = \log \left( 1 + \frac{1}{2} \|P_{Y|X}(\cdot|1) - P_{Y|X}(\cdot|0)\| \right),$$

where  $\|\cdot\|_1$  is the  $L_1$  distance. The term inside the log is twice the probability of success in binary hypothesis testing, which sheds light on why  $I_\infty(X; Y)$  arises as maximal leakage. We evaluate  $\mathcal{L}(X \rightarrow Y)$  for some special cases.

*Example 1:* If  $X \sim \text{Ber}(q)$ ,  $0 < q < 1$ , and  $Y$  is the output of a BSC with parameter  $p$ ,  $0 \leq p \leq 1/2$ , then  $\mathcal{L}(X \rightarrow Y) = \log(2(1-p))$ .

*Example 2:* If  $X \sim \text{Ber}(q)$ ,  $0 < q < 1$ , and  $Y$  is the output of a BEC with parameter  $\epsilon$ ,  $0 \leq \epsilon \leq 1$ , then  $\mathcal{L}(X \rightarrow Y) = \log(2 - \epsilon)$ , and  $\mathcal{L}(Y \rightarrow X) = \log 2$ .

*Example 3:* For any deterministic law  $P_{Y|X}$ ,  $\mathcal{L}(X \rightarrow Y) = \log |\{y : P_Y(y) > 0\}|$ .

The following corollary summarizes some useful properties of  $\mathcal{L}(X \rightarrow Y)$ .

**Corollary 2:** For any joint distribution  $P_{XY}$  on finite alphabets  $\mathcal{X}$  and  $\mathcal{Y}$ ,

- 1) (*Data Processing Inequality*) If the Markov chain  $X - Y - Z$  holds for a discrete random variable  $Z$ , then  $\mathcal{L}(X \rightarrow Z) \leq \min\{\mathcal{L}(X \rightarrow Y), \mathcal{L}(Y \rightarrow Z)\}$ .
- 2)  $\mathcal{L}(X \rightarrow X) = H_0(X) = \log |\{x : P_X(x) > 0\}|$ .
- 3)  $\mathcal{L}(X \rightarrow Y) \leq \min\{\log |\mathcal{X}|, \log |\mathcal{Y}|\}$ .
- 4)  $\mathcal{L}(X \rightarrow Y) \geq I(X; Y)$ .
- 5)  $\mathcal{L}(X \rightarrow Y) = 0$  iff  $X$  and  $Y$  are independent.
- 6)  $\mathcal{L}(X \rightarrow Y)$  is not symmetric in  $X$  and  $Y$ .
- 7) (*Additivity*) If  $\{(X_i, Y_i)\}_{i=1}^\ell$  are mutually independent, then

$$\mathcal{L}(X_1^\ell \rightarrow Y_1^\ell) = \sum_{i=1}^{\ell} \mathcal{L}(X_i \rightarrow Y_i).$$

- 8)  $\exp\{\mathcal{L}(X \rightarrow Y)\}$  is convex in  $P_{Y|X}$  for fixed  $P_X$ .

*Proof:* Properties 1) through 4), and 7) are shown for  $I_\infty(X; Y)$  [19]. 5) follows from the definition and 4). 6) is clear and is illustrated in Example 2. 8) follows from the fact that, for each  $y \in \mathcal{Y}$ ,  $\max_x P_{Y|X}(y|x)$  is convex in  $P_{Y|X}$ . ■

Note that properties 1), 5), and 7) can be regarded as axiomatic for a leakage measure. Property 4) shows that a small maximal leakage is a more stringent requirement than a small mutual information. Property 8) shows that minimizing maximal leakage, for a fixed  $P_X$ , amounts to minimizing a convex function.

#### IV. MAXIMAL LEAKAGE VARIATIONS

We show the robustness of maximal leakage, by proving that variations on its definition yield the same quantity.

##### A. $k$ -maximal leakage

We allow the adversary several guesses, as arises in some practical situations discussed in the introduction.

*Definition 2 ( $k$ -Maximal Leakage):* Given a joint distribution  $P_{XY}$  on finite alphabets  $\mathcal{X}$  and  $\mathcal{Y}$ , and a positive integer  $k$ , the  $k$ -maximal leakage from  $X$  to  $Y$  is defined as

$$\mathcal{L}^{(k)}(X \rightarrow Y) = \sup_{U-X-Y-(\hat{U}_i)_{i=1}^k} \log \frac{\Pr \left( \bigvee_{i=1}^k U = \hat{U}_i \right)}{\max_{\substack{S \subseteq \mathcal{U} \\ |S| \leq k}} P_U(S)}.$$

The following lemma establishes the equivalence between maximal leakage and  $k$ -maximal leakage.

**Lemma 1:** For any joint distribution  $P_{XY}$  on finite alphabets  $\mathcal{X}$  and  $\mathcal{Y}$ , and any  $k \in \mathbb{N}$ ,

$$\mathcal{L}^{(k)}(X \rightarrow Y) = \mathcal{L}(X \rightarrow Y).$$

*Proof:*

To show  $\mathcal{L}^{(k)}(X \rightarrow Y) \geq \mathcal{L}(X \rightarrow Y)$ , for any  $P_{U|X}$ , we construct  $P_{V|X}$  such that  $\mathcal{L}^{(k)}(X \rightarrow Y)[V] =$

$\mathcal{L}(X \rightarrow Y)[U]$ . In particular, for a given  $P_{U|X}$  and associated alphabet  $\mathcal{U}$ , let

$$\mathcal{V} = \bigcup_{u \in \mathcal{U}} \{(u, 1), (u, 2), \dots, (u, k)\},$$

and  $P_{V|X}(v|x) = P_{V|X}((a_v, b_v)|x) = P_{U|X}(a_v|x)/k$ .

Then, observing  $Y$ , the probability of guessing  $V$  correctly with  $k$  guesses is:

$$\begin{aligned} & \sup_{X \rightarrow Y - (\hat{V}_i)_{i=1}^k} \Pr(V = \hat{V}_1 \vee \dots \vee V = \hat{V}_k) \\ &= \sum_{y \in \mathcal{Y}} \max_{v_1, v_2, \dots, v_k} \sum_{i=1}^k \sum_{x \in \mathcal{X}} P_X(x) P_{V|X}(v_i|x) P_{Y|X}(y|x) \\ &= \sum_{y \in \mathcal{Y}} \sum_{i=1}^k \max_{v_i \neq v_1, \dots, v_{i-1}} \sum_{x \in \mathcal{X}} P_X(x) P_{V|X}(v_i|x) P_{Y|X}(y|x) \\ &\stackrel{(a)}{=} \sum_{y \in \mathcal{Y}} \max_u \sum_{x \in \mathcal{X}} P_X(x) P_{U|X}(u|x) P_{Y|X}(y|x), \end{aligned} \quad (5)$$

where (a) follows by setting  $v_i = (u^*, i)$ , where

$$u^* = \operatorname{argmax}_{u \in \mathcal{U}} \sum_{x \in \mathcal{X}} P_X(x) P_{U|X}(u|x) P_{Y|X}(y|x).$$

Now, note that (5) is simply the probability of guessing  $U$  correctly with a single guess after observing  $Y$ . A similar argument shows that, with no  $Y$  observation, the probability of guessing  $V$  correctly with  $k$  guesses is equal to the probability of guessing  $U$  correctly with a single guess, hence  $\mathcal{L}^{(k)}(X \rightarrow Y)[V] = \mathcal{L}(X \rightarrow Y)[U]$ , which establishes  $\mathcal{L}^{(k)}(X \rightarrow Y) \geq \mathcal{L}(X \rightarrow Y)$ .

We still need to show  $\mathcal{L}(X \rightarrow Y) \geq \mathcal{L}^{(k)}(X \rightarrow Y)$ . For any  $P_{V|X}$ , we construct  $P_{U|X}$  such that  $\mathcal{L}(X \rightarrow Y)[U] = \mathcal{L}^{(k)}(X \rightarrow Y)[V]$ . So let  $P_{V|X}$  be given, with associated alphabet  $\mathcal{V}$ , and let  $\ell \triangleq |\mathcal{V}| \geq k$ . Now, let

$$\mathcal{U} = \{S \subset \mathcal{V} : |S| = k\},$$

$$\text{and } p_{U|X}(u|x) = c \sum_{v \in u} p_{V|X}(v|x),$$

where  $c = 1/\binom{\ell-1}{k-1}$ . Then, observing  $Y$ , the probability of guessing  $U$  correctly with a single guess is

$$\begin{aligned} & \sup_{X \rightarrow Y - \hat{U}} \Pr(U = \hat{U}) \\ &= \sum_{y \in \mathcal{Y}} \max_{u \in \mathcal{U}} \sum_{x \in \mathcal{X}} P_X(x) P_{U|X}(u|x) P_{Y|X}(y|x) \\ &= \sum_{y \in \mathcal{Y}} \max_{u \in \mathcal{U}} \sum_{x \in \mathcal{X}} P_X(x) \sum_{v \in u} P_{V|X}(v|x) P_{Y|X}(y|x) c \\ &= c \sum_{y \in \mathcal{Y}} \max_{v_1, v_2, \dots, v_k} \sum_{x \in \mathcal{X}} \sum_{i=1}^k P_X(x) P_{V|X}(v_i|x) P_{Y|X}(y|x), \end{aligned}$$

which is the probability, normalized by  $c$ , of guessing  $V$  correctly with  $k$  guesses after observing  $Y$ . A similar argument shows that, with no  $Y$  observation, the probability of guessing  $U$  correctly with a single guess is equal to the probability, normalized by  $c$ , of guessing  $V$  correctly with  $k$  guesses, hence  $\mathcal{L}(X \rightarrow Y)[V] = \mathcal{L}^{(k)}(X \rightarrow Y)[U]$ , which establishes  $\mathcal{L}(X \rightarrow Y) \geq \mathcal{L}^{(k)}(X \rightarrow Y)$ . ■

### B. Maximal locational leakage

For locational leakage, the adversary only needs to generate a guess that is within a certain distance of the true function value. The term ‘‘locational’’ is motivated by the scenario in which the variable of interest  $U$  is a geographical location.

*Definition 3 (Maximal Locational Leakage):* Given a joint distribution  $P_{XY}$  on finite alphabets  $\mathcal{X}$  and  $\mathcal{Y}$ , and a metric space  $\mathcal{U}$  (with its associated Borel  $\sigma$ -field), the maximal locational leakage from  $X$  to  $Y$  is defined as

$$\mathcal{L}_{\mathcal{U}}(X \rightarrow Y) = \sup_{U: X \rightarrow Y} \log \frac{\sup_{\hat{u}(\cdot)} \Pr(U \in B(\hat{u}(Y)))}{\sup_{\hat{u}} \Pr(U \in B(\hat{u}))}, \quad (6)$$

where  $B(u)$  denotes the closed unit ball centered at  $u \in \mathcal{U}$ .

**Lemma 2:** For any joint distribution  $P_{XY}$  on finite alphabets  $\mathcal{X}$  and  $\mathcal{Y}$ , and any metric space  $\mathcal{U}$ ,

$$\mathcal{L}_{\mathcal{U}}(X \rightarrow Y) \leq \mathcal{L}(X \rightarrow Y),$$

with equality if  $\mathcal{U}$  has an infinitely countable subset  $S$ , such that no pair of its elements can be contained in a single unit ball.

*Proof:*

Consider any  $U$  and  $\hat{u}(Y)$  in the maximization of (6):

$$\begin{aligned} & \Pr(U \in B(\hat{u}(Y))) \\ & \leq \sum_{y \in \mathcal{Y}} \sup_{u \in \mathcal{U}} P(U \in B(u), Y = y) \\ & = \sum_{y \in \mathcal{Y}} \sup_{u \in \mathcal{U}} \sum_{x \in \mathcal{X}} P(U \in B(u), X = x, Y = y) \\ & = \sum_{y \in \mathcal{Y}} \sup_{u \in \mathcal{U}} \sum_{x \in \mathcal{X}} P(U \in B(u)) \cdot \\ & \quad P(X = x | U \in B(u)) P_{Y|X}(y|x) \\ & \leq \sum_{y \in \mathcal{Y}} \sup_{u \in \mathcal{U}} P(U \in B(u)) \sup_{x \in \mathcal{X}} p_{Y|X}(y|x) \\ & = \left[ \sum_{y \in \mathcal{Y}} \sup_{x \in \mathcal{X}} p_{Y|X}(y|x) \right] \sup_{u \in \mathcal{U}} P(U \in B(u)). \end{aligned}$$

Therefore,

$$\mathcal{L}_{\mathcal{U}}(X \rightarrow Y) \leq \log \sum_{y \in \mathcal{Y}} \sup_{x \in \mathcal{X}} p_{Y|X}(y|x) = \mathcal{L}(X \rightarrow Y).$$

If  $\mathcal{U}$  satisfies the lemma condition (e.g.,  $\mathcal{U}$  is unbounded), then exact guessing of discrete functions can be simulated by choosing  $S$  to be the support of  $U$ . Hence  $\mathcal{L}_{\mathcal{U}}(X \rightarrow Y) \geq \mathcal{L}(X \rightarrow Y)$ , which implies the equality. ■

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